



MALLIAVIN CALCULUS WITH APPLICATIONS TO FINANCE

Keywords: Stochastic Calculus, Probability Theory, Finance

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1 Context Setting

Consider a market consisting of a riskless asset S_0 with

$$\text{riskless asset} \quad \begin{cases} dS_0(t) = \rho(t)S_0(t) dt \\ S_0(0) = 1 \end{cases} \quad (1)$$

and a risky asset S_1 satisfying

$$\text{risky asset} \quad \begin{cases} dS_1(t) = \mu(t)S_1(t) dt + \sigma(t)S_1(t) dW(t) \\ S_1(0) > 0 \end{cases} \quad (2)$$

where $\rho(t)$, $\mu(t)$, and $\sigma(t) \neq 0$ are \mathbf{F} -adapted processes satisfying the following condition

$$\mathbf{E} \left[\int_0^T (|\rho(t)| + |\mu(t)| + \sigma^2(t)) dt \right] < \infty$$

Let $\theta_0(t)$ and $\theta_1(t)$ denote the number of units of $S_0(t)$ and $S_1(t)$, respectively. Then the value of the portfolio $\theta = (\theta_0, \theta_1)$ is $V^\theta = \theta_0 S_0 + \theta_1 S_1$.

We also suppose that the portfolio is self-financing, i.e.,

$$dV^\theta(t) = \theta_0(t)dS_0(t) + \theta_1(t)dS_1(t) \quad (3)$$

Substituting

$$\theta_0(t) = \frac{V^\theta(t) - \theta_1(t)S_1(t)}{S_0(t)}$$

¹This research was supported by the São Paulo Research Foundation (FAPESP). Process number: 2023/08064-0.

into (3) and using (1) we have

$$dV^\theta = \rho(t)(V^\theta(t) - \theta_1(t)S_1(t))dt + \theta_1(t)dS_1 \quad (4)$$

Replacing (2),

$$dV^\theta = [\rho(t)V^\theta(t) + (\mu(t) - \rho(t))\theta_1(t)S_1(t)]dt + \sigma(t)\theta_1(t)S_1(t)dW(t) \quad (5)$$

Our goal is to find a replicating (hedging) portfolio

$$V^\theta(T) = F, \quad \mathbf{P}\text{-a.s.} \quad (6)$$

where F is \mathcal{F}_T -measurable. For an European call, for example, $F = \max\{S_1 - K, 0\} = (S_1 - K)^+$.

How much do we need to invest at time $t = 0$ and which portfolio $\theta(t)$ should we use? Are V^θ and θ unique?

We consider $(V^\theta(t), \theta_1(t))$ an \mathbf{F} -adapted process. The equations (4) and (6) form a **backward stochastic differential equation** (BSDE). To find an explicit solution, we can change the measure and apply Clark-Ocone.

Define

$$u(t) = \frac{\mu(t) - \rho(t)}{\sigma(t)} \quad (7)$$

Using the change of measure as in the last section, we can write

$$\begin{aligned} dV^\theta &= [\rho(t)V^\theta(t) + (\mu(t) - \rho(t))\theta_1(t)S_1(t)]dt + \sigma(t)\theta_1(t)S_1(t)d\tilde{W}(t) \\ &\quad - \sigma(t)\theta_1(t)S_1(t)\sigma^{-1}(t)(\mu(t) - \rho(t))dt \\ &= \rho(t)V^\theta(t)dt + \sigma(t)\theta_1(t)S_1(t)d\tilde{W}(t) \end{aligned} \quad (8)$$

Let

$$U^\theta(t) = e^{-\int_0^t \rho(s) ds} V^\theta(t)$$

Then using (8),

$$dU^\theta(t) = e^{-\int_0^t \rho(s) ds} \sigma(t)\theta_1(t)S_1(t) d\tilde{W}(t)$$

or, equivalently,

$$e^{-\int_0^t \rho(s) ds} V^\theta(T) = V^\theta(0) + \int_0^T e^{-\int_0^t \rho(s) ds} \sigma(t)\theta_1(t)S_1(t) d\tilde{W}(t) \quad (9)$$

Let us recall the following result.

Theorem 1.1 (Clark-Ocone for $L^2(\mathbf{P})$ Under Change of Measure). Suppose that $F \in L^2(\mathbf{P})$ is \mathcal{F}_T -measurable, and that the following conditions are met

1. $\mathbf{E}_Q[|F|] < \infty$;
2. $\mathbf{E}_Q\left[\int_0^T |D_t F|^2 dt\right] < \infty$;
3. $\mathbf{E}_Q\left[|F| \int_0^T \left(\int_0^T D_t u(s) dW(s) + \int_0^T u(s)D_t u(s) ds\right)^2 dt\right] < \infty$.

Then

$$F = \mathbf{E}_Q[F] + \int_0^T \mathbf{E}_Q \left[\left(D_t F - F \int_t^T D_t u(s) d\tilde{W}(s) \right) \middle| \mathcal{F}_t \right] d\tilde{W}(t)$$

where $\tilde{W}(t)$ is a Brownian motion under the measure Q and $D_t F \in \mathcal{G}^*$ is the Hida-Malliavin derivative.

Proof. Analogous to the case for $\mathbf{D}_{1,2}$. See, e.g. [Oku10]. □

Applying the generalized Clark-Ocone formula to

$$G = e^{-\int_0^t \rho(s) ds} F$$

we have

$$G = \mathbf{E}_Q[G] + \int_0^T \mathbf{E}_Q \left[\left(D_t G - G \int_t^T D_t u(s) d\tilde{W}(s) \right) \middle| \mathcal{F}_t \right] d\tilde{W}(t) \quad (10)$$

Comparing (9) with (10), we have $V^\theta(0) = \mathbf{E}_Q[G]$ by uniqueness, and the replicating portfolio is given by

$$\theta_1(t) = e^{-\int_0^t \rho(s) ds} \sigma^{-1}(t) S_1^{-1}(t) \mathbf{E}_Q \left[\left(D_t G - G \int_t^T D_t u(s) d\tilde{W}(s) \right) \middle| \mathcal{F}_t \right] \quad (11)$$

In particular, if ρ and μ are constants, and $\sigma(t) = \sigma \neq 0$, then

$$u(t) = u = \frac{\mu - \rho}{\sigma}$$

is also constant, whence $D_t u = 0$. Then the equation (11) simplifies to

$$\theta_1(t) = e^{\rho(t-T)} \sigma^{-1} S_1^{-1}(t) \mathbf{E}_Q[D_t F \mid \mathcal{F}_t] \quad (12)$$

In this presentation, we consider a digital option, which has a payoff at maturity

$$F = \mathbf{1}_{[K, \infty)}(W(T))$$

where K is the exercise price. We aim to compute the conditional expectation $\mathbf{E}_Q[D_t F \mid \mathcal{F}_t]$.

2 Necessary Results

To do that, we need the following concept.

Definition 2.1 (Donsker delta function). Let $Y : \Omega \rightarrow \mathbf{R}$, $Y \in \mathcal{G}^*$. The continuous function

$$\delta_Y(\cdot) : \mathcal{R} \rightarrow \mathcal{G}^*$$

is the **Donsker delta function** of Y if it has the property that

$$\int_{\mathbf{R}} f(y) \delta_Y(y) dy = f(Y) \quad \text{a.s.}$$

for all measurable $f : \mathbf{R} \rightarrow \mathbf{R}$ such that the integral converges.

Theorem 2.1. Suppose that

1. $\alpha : [0, T] \rightarrow \mathbf{R}^n$ is a deterministic function such that $\|\alpha\|^2 = \int_0^T \alpha^2(s) ds < \infty$.
2. $\varphi : [0, T] \rightarrow \mathbf{R}^{n \times n}$ is a deterministic function such that $\|\varphi\|^2 = \sum_{i,j=1}^n \int_0^T \varphi_{ij}^2(s) ds < \infty$.
3. $f : \mathbf{R}^n \rightarrow \mathbf{R}$ is bounded.

Define, for $t \in [0, T]$,

$$Y(t) = \int_0^t \varphi(s) dB(s) + \int_0^t \varphi(s) \alpha(s) ds$$

Then

$$f(Y(T)) = V_0 + \int_0^T u(t, \omega) \diamond (\alpha(t) + W(t)) dt$$

where $|A| = \det A$ and A is the inverse of the covariance matrix of Y ,

$$V_0 = (2\pi)^{-n/2} \sqrt{|A|} \int_{\mathbf{R}^n} f(y) \exp\left(-\frac{1}{2} y^T A y\right) dy$$

and

$$u(t, \omega) = (2\pi)^{-n/2} \sqrt{|A|} \int_{\mathbf{R}^n} f(y) \exp^\diamond\left(-\frac{1}{2} (y - Y(t))^T \diamond A (y - Y(t))\right) \diamond ((y - Y(t))^T A \varphi(t)) dy$$

Proof. Refer to [AØU01, Theorem 4.4]. □

The next result is a simpler version of [HØ03, Lemma 3.21].

Theorem 2.2. Let $\diamond_{\mathbf{P}}$ and $\diamond_{\mathbf{Q}}$ denote the Wick product for the probability measures \mathbf{P} and \mathbf{Q} respectively, and u as in the Girsanov Theorem. Then, $F \diamond_{\mathbf{P}} G = F \diamond_{\mathbf{Q}} G$.

Proof. Let $F = \exp^\diamond\left(\int_0^\infty f(t) dW(t)\right)$ and $G = \exp^\diamond\left(\int_0^\infty g(t) dW(t)\right)$. Then,

$$F \diamond_{\mathbf{P}} G = \exp\left(\int_0^\infty (f(s) + g(s)) dW(s) - \frac{1}{2} \|f + g\|^2\right)$$

Applying Girsanov to F and G ,

$$F = \exp^\diamond\left(\int_0^\infty f(s) d\tilde{W}(s) - \langle f, u \rangle_{L^2}\right), \quad \text{and} \quad G = \exp^\diamond\left(\int_0^\infty g(s) d\tilde{W}(s) - \langle g, u \rangle_{L^2}\right)$$

Computing the product,

$$\begin{aligned} F \diamond_{\mathbf{Q}} G &= \exp\left(\int_0^\infty (f(s) + g(s)) d\tilde{W}(s) - \frac{1}{2} \|f + g\|^2 - \langle f + g, u \rangle_{L^2}\right) \\ &= \exp\left(\int_0^\infty (f(s) + g(s)) dW(s) - \frac{1}{2} \|f + g\|^2\right) \end{aligned}$$

Hence, $F \diamond_{\mathbf{P}} G = F \diamond_{\mathbf{Q}} G$ for exponential functions. By density, we have the result. □

3 Main Result

Theorem 3.1. Suppose that ρ is constant, and $u(t)$, as defined in (7), is deterministic and satisfying $\mathbf{E}[u^2(t)] < \infty$. Then the replicating portfolio for hedging $\mathbf{1}_{[K, \infty)}(W(T))$ is

$$\theta_1(t) = e^{-\rho(T-t)} (2\pi(T-t))^{-1/2} \sigma^{-1}(t) S_1^{-1}(t) \exp\left(-\frac{(K-W(t))^2}{2(T-t)}\right) \quad (13)$$

Proof. First, notice that $F = \mathbf{1}_{[K, \infty)}(W(T)) \in L^2(\mathbf{P})$. Thus, we can use (12).

Now we compute $\mathbf{E}_Q[D_t F \mid \mathcal{F}_t]$ using the Donsker delta function by taking $f(y) = \mathbf{1}_{[K, \infty)}(y)$, and $Y(T) = W(T)$. By the Theorem 2.1,

$$\mathbf{1}_{[K, \infty)}(W(T)) = \int_K^\infty (2\pi T)^{-1/2} \exp^\diamond\left(-\frac{(y-W(T))^{\diamond 2}}{2T}\right) dy$$

By the Chain Rule for the Wick product,

$$\begin{aligned} D_t(\mathbf{1}_{[K, \infty)}(W(T))) &= \int_K^\infty (2\pi T)^{-1/2} \exp^\diamond\left(-\frac{(y-W(T))^{\diamond 2}}{2T}\right) \diamond \frac{(y-W(T))}{2T} dy \\ &= (2\pi T)^{-1/2} \exp^\diamond\left(-\frac{(K-W(T))^{\diamond 2}}{2T}\right) \end{aligned}$$

Denoting by $\hat{\diamond}$ the Wick product with respect to the probability measure Q , then since $\hat{\diamond} = \diamond$ (Theorem 2.2), we have

$$\begin{aligned} \mathbf{E}[D_t(\mathbf{1}_{[K, \infty)}(W(T))) \mid \mathcal{F}_t] &= \mathbf{E}_Q \left[(2\pi T)^{-1/2} \exp^{\hat{\diamond}}\left(-\frac{(K-W(T))^{\hat{\diamond} 2}}{2T}\right) \mid \mathcal{F}_t \right] \\ &= (2\pi T)^{-1/2} \mathbf{E}_Q \left[\exp^{\hat{\diamond}}\left(-\frac{(K-\tilde{W}(T) + \int_0^T u(s) ds)^{\hat{\diamond} 2}}{2T}\right) \mid \mathcal{F}_t \right] \\ &= (2\pi T)^{-1/2} \exp\left(-\frac{(K-W(t))^2}{2(T-t)}\right) \end{aligned}$$

□

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